

$$\begin{aligned}
 &= E \left[ e^{t(x_1+x_2)} \right] = E(e^{tx_1}) \cdot E(e^{tx_2}) \\
 &= (q+pe^t)^{n_1} \cdot (q+pe^t)^{n_2} \\
 &= (q+pe^t)^{n_1+n_2}
 \end{aligned}$$

This mgf also have binomial distribution with parameter  $(n_1+n_2, p)$

Q:- Let  $X_1, X_2$  and  $X_3$  are the independent poisson random variables having mean 2, 1 and 4

- Find mgf of  $Y = X_1 + X_2 + X_3$
- How  $Y$  is distributed
- compute  $P(3 \leq Y \leq 9)$

$$\begin{aligned}
 M_Y(t) &= E(e^{ty}) = E \left[ e^{t(X_1+X_2+X_3)} \right] \\
 &= E(e^{tx_1}) \cdot E(e^{tx_2}) \cdot E(e^{tx_3}) \\
 &= e^{\lambda_1(t-1)} \cdot e^{\lambda_2(t-1)} \cdot e^{\lambda_3(t-1)} \\
 &= e^{2(t-1)} \cdot e^{1(t-1)} \cdot e^{4(t-1)} \\
 &= e^{7(t-1)}
 \end{aligned}$$

Distribution of  $Y$  is also poisson with parameter  $\lambda = 7$ .

$$P(3 \leq Y \leq 9) = \frac{e^{-7} 7^3}{3!} + \frac{e^{-7} 7^4}{4!} + \frac{e^{-7} 7^5}{5!} +$$

$$\frac{e^{-7} 7^6}{6!} + \frac{e^{-7} 7^7}{7!} + \frac{e^{-7} 7^8}{8!} + \frac{e^{-7} 7^9}{9!}$$

$$= 0.0521 + 0.0912 + 0.1277 + 0.1490 +$$

$$0.1490 + 0.1303 + 0.1014$$

$$= 0.8007.$$

Q:- Let XYZ are jointly continuous random variables with joint Pdf

$$f(x,y,z) = \frac{1}{3}(x+2y+3z) \quad 0 \leq x,y,z \leq 1$$

Find joint Pdf of X and Y.

Joint Pdf of x and y = ?

$$f(x,y) = \int_0^1 \frac{1}{3}(x+2y+3z) dz$$

$$= \int_0^1 \frac{x}{3} dz + \frac{2}{3} \int_0^1 y dz + \frac{3}{3} \int_0^1 z dz$$

$$= \frac{x}{3} \int_0^1 dz + \frac{2y}{3} \int_0^1 dz + \frac{3}{3} \int_0^1 z dz$$

$$= \frac{x}{3} z \Big|_0^1 + \frac{2y}{3} z \Big|_0^1 + \frac{3z^2}{3 \cdot 2} \Big|_0^1$$

$$= \frac{x}{3} + \frac{2y}{3} + \frac{1}{2}$$

$$= \frac{1}{3} (x + 2y + \frac{3}{2})$$

Q:- Let  $x, y, z$  are 3 independent random variables with  $x \sim N(\mu, \sigma^2)$  and  $y$  and  $z \sim U(0, 2)$ . We also have expected value of  $(x^2y + xy^2z) = 13$  and expected value of  $(xy^2 + x^2z) = 14$ . Find  $\mu$  and  $\sigma$ .

Mean and variance of uniform distribution

$$E(y) = E(z) = \frac{a+b}{2} = \frac{2}{2} = 1$$

$$V(y) = V(z) = \frac{(b-a)^2}{12} = \frac{4}{12} = \frac{1}{3}$$

Now

$$E(x^2y + xy^2z) = 13$$

$$E(x^2y) + E(xy^2z) = 13$$

$$[E(x^2) \cdot E(y)] + [E(x) \cdot E(y) \cdot E(z)] = 13$$

$$[(\sigma^2 + \mu^2) \cdot 1] + \mu \cdot 1 = 13$$

$$\sigma^2 + \mu^2 + \mu = 13 \quad \text{--- (A)}$$

$$\because E(x) = \mu$$

$$E(x^2) = [E(x)]^2 + V(x) = \mu^2 + \sigma^2$$

Again

$$E(xy^2 + x^2z) = 14$$

$$[E(x) \cdot E(y^2)] + [E(x^2) \cdot E(z)] = 14$$

$$\begin{aligned}\because E(y^2) &= V(y) + [E(y)]^2 \\ &= \frac{1}{3} + 1 = \frac{4}{3}\end{aligned}$$

$$\mu \cdot \frac{4}{3} + (\mu^2 + \sigma^2) = 14$$

$$\frac{4\mu}{3} + \mu^2 + \sigma^2 = 14$$

$$4\mu + 3\mu^2 + 3\sigma^2 = 42 \quad \text{--- (B)}$$

Solving (A) and (B)

$$\begin{aligned}\sigma^2 + \mu^2 + \mu &= 13 \\ 3\sigma^2 + 3\mu^2 + 4\mu &= 42 \quad \Rightarrow \quad \begin{aligned}3\sigma^2 + 3\mu^2 + 3\mu &= 39 \\ \underline{3\sigma^2 + 3\mu^2 + 4\mu} &= \underline{42}\end{aligned}\end{aligned}$$

$$-\mu = -3$$

$$\mu = 3$$

Putting value of  $\mu = 3$  in eq (A)

$$\sigma^2 + 9 + 3 = 13$$

$$\sigma^2 + 12 = 13$$

$$\sigma^2 = 1, \quad \sigma = 1.$$

## Law of large number:-

Law of large number is used to established the consistency of results

Detail:- In probability theory, law of large number is a theorem that describe the results of performing some experiment a large number of times. According to this law the average of results obtained from a large number of trials should be closed to expected value of results and will tends to become closer as more trials are performed

∴ Law of large no. also called law of average

These are two types of law of large number

1) weak law of large number (WLLN):-

2) strong law of large number (SLLN):-

1) Weak law of large numbers:-

Let  $X_1, X_2, \dots, X_n$

are independent trials with expected value of  $X_i$   $[E(X_i)] = \mu$  and finite variance  $\sigma^2$ . Let  $S_n = \text{Sum of indep. trials} = X_1 + X_2 + \dots + X_n$ . Then for any arbitrary value  $\epsilon > 0$  the WLLN states that the sample average  $\left(\frac{S_n}{n}\right)$  converges in probability towards the expected value.

Mathematically:-

$$P\left[\left|\frac{S_n}{n} - \mu\right| > \epsilon\right] = 0 \quad \text{as } n \rightarrow \infty$$

$$\because \frac{S_n}{n} = \bar{x} \quad P\left[\left|\frac{S_n}{n} - \mu\right| \leq \epsilon\right] = 1 \quad \text{as } n \rightarrow \infty$$

$$\because \mu = E(\bar{x})$$

$$\bar{x} \xrightarrow{P} \mu$$

Here  $P$  means "converges in probability".

2) Strong law of large numbers:-

Let  $X_1, X_2, \dots, X_n$  are

the independent trials with  $E(X_i) = \mu$  and finite variance  $\sigma^2$ .  $S_n = X_1 + X_2 + \dots + X_n$

then for any no.  $\epsilon > 0$  the strong law of large numbers states that the sample average converges almost surely to the expected value

Mathematically :-

$$\frac{S_n}{n} \xrightarrow{\text{a.s.}} \mu$$

$\therefore$  a.s. = almost surely  
(یقینی)

Q:- State and prove law of large numbers.

Statement :-

Let  $S_n$  be a sequence of independent random variables with finite  $\mu$  and finite variance  $\sigma^2$ . Then for any number  $\epsilon > 0$

$$P\left\{\left[\frac{S_n}{n} - E\left(\frac{S_n}{n}\right)\right] > \epsilon\right\} \xrightarrow{P} 0 \quad n \rightarrow \infty$$

$$\lim_{n \rightarrow \infty} P\left\{\left[\frac{S_n}{n} - E\left(\frac{S_n}{n}\right)\right] > \epsilon\right\} \xrightarrow{P} 0$$

Proof :-

$$S_n = x_1 + x_2 + \dots + x_n$$

then  $\frac{S_n}{n} = \bar{x}_n$

$$E(\bar{X}_n) = E\left(\frac{S_n}{n}\right)$$

$$= E\left(\frac{X_1 + X_2 + \dots + X_n}{n}\right)$$

$$= \frac{1}{n} (E(X_1) + E(X_2) + \dots + E(X_n))$$

$$= \frac{1}{n} (\mu + \mu + \dots + \mu)$$

$$= \frac{n\mu}{n} = \mu$$

$$V(\bar{X}_n) = V\left(\frac{S_n}{n}\right) = \frac{1}{n^2} V(S_n)$$

$$= \frac{1}{n^2} \{V(X_1) + V(X_2) + \dots + V(X_n)\}$$

$$= \frac{1}{n^2} \{\sigma^2 + \sigma^2 + \dots + \sigma^2\} = \frac{n\sigma^2}{n^2} = \frac{\sigma^2}{n}$$

According to Chebyshev's inequality

$$P\left\{|\bar{X}_n - E(\bar{X}_n)| > \epsilon\right\} \leq \frac{V(\bar{X}_n)}{\epsilon^2}$$

$$P\left\{|\bar{X}_n - \mu| > \epsilon\right\} \leq \frac{\sigma^2}{n\epsilon^2}$$

Applying limit on both sides

$$\lim_{n \rightarrow \infty} P\left\{|\bar{X}_n - \mu| > \epsilon\right\} \leq \lim_{n \rightarrow \infty} \frac{\sigma^2}{n\epsilon^2}$$

$$\lim_{n \rightarrow \infty} P\left\{|\bar{X}_n - \mu| > \epsilon\right\} \leq 0$$



## Classical Central Limit Theorem:-

As sample size increases the sampling distribution of sample mean approaches to the normal distribution with mean as the population and standard deviation is the population standard deviation by square root of  $n$  (sample size).

$$\bar{x} \sim N(\mu, \sigma/\sqrt{n})$$

$$\bar{x} - \mu \sim N(0, \sigma/\sqrt{n})$$

$$\frac{\bar{x} - \mu}{\sigma/\sqrt{n}} \sim N(0, 1)$$

## Central Limit Theorem and Properties:-

Let  $X_1, X_2, \dots, X_n$  be

a sequence of  $n$  independent and identically distributed random variables having finite population mean  $\mu$  and variance  $\sigma^2$ . Then the CLT states that as sample size increase the distribution of sample statistic of the random variables approaches to normal distribution

with mean  $\mu_{\bar{x}} = \mu$  and variance

$$\sigma_{\bar{x}}^2 = \sigma^2/n.$$

If  $\uparrow$  the sample average  
or sample mean  $\downarrow$  we define  $\bar{x}_n$  be

$\bar{X}_n \sim N(\mu, \sigma^2/n)$ . In a similar way

$$Z_n = \frac{\bar{X}_n - \mu}{\sigma/\sqrt{n}}$$

Proof:-

We know that moment generating function (mgf) of standard normal distribution

$$M_{Z_n}(t) = e^{t^2/2}$$

Let  $M_{Z_n}(t)$  denote the mgf of  $Z_n$  and we have to show when  $n \rightarrow \infty$

$$\because Z_n = \frac{\bar{X}_n - \mu}{\sigma/\sqrt{n}}$$

$$M_{Z_n}(t) \rightarrow M_X(t) = e^{t^2/2}$$

By definition:-

$$\begin{aligned} M_{Z_n}(t) &= E[e^{t(Z_n)}] \\ &= E\left[e^{\frac{t(\bar{X}_n - \mu)\sqrt{n}}{\sigma}}\right] \\ &= E\left[e^{\frac{t \cdot (\sum_{i=1}^n X_i - n\mu)\sqrt{n}}{\sigma}}\right] \\ &= E\left[e^{\frac{t}{n} \sum_{i=1}^n (X_i - \mu)\sqrt{n}}\right] \\ &= E\left[\prod_{i=1}^n e^{\frac{t}{\sqrt{n}} \frac{(X_i - \mu)}{\sigma}}\right] \quad \text{--- (1)} \end{aligned}$$

$\because$  Since  $X_i$ 's are independent

Let  $Y_i = \frac{X_i - \mu}{\sigma}$  then

$M_{Y_i}(t)$  is the mgf of  $X_i$  and

$Y_i$  has the same distribution as  $X_i$ . Now (1) becomes as

$$M_{Z_n}(t) = \prod_{i=1}^n E\left(e^{t \cdot Y_i / \sqrt{n}}\right)$$

$$= \prod_{i=1}^n M_{Y_i}(t/\sqrt{n})$$

$$= \left[M_{Y_i}(t/\sqrt{n})\right]^n \quad \text{--- (2)}$$

$\therefore$  By using property of mgf when variables are iid then product changes into power

we know that

$$M_X(t) = M_0(t) + \frac{t}{1!} M'_1(t) + \frac{t^2}{2!} M''_2(t) \\ + \frac{t^3}{3!} M'''_3(t) + \dots$$

By using this relation

$$M_{Y_i}(t/\sqrt{n}) = 1 + \frac{t}{\sqrt{n}} \left(\frac{\mu_1}{\sigma}\right) + \frac{\mu_2}{2\sigma^2} \left(\frac{t}{\sqrt{n}}\right)^2 \\ + \frac{\mu_3}{3! \sigma^3} \left(\frac{t}{\sqrt{n}}\right)^3 + \dots$$

Since all odd order moments about mean are zero so

$$= 1 + 0 + \frac{\sigma^2}{2! \sigma^2} \cdot \left(\frac{t}{\sqrt{n}}\right)^2 + \dots$$

$$= 1 + \frac{t^2}{2n} + \dots$$

$$= 1 + P/n$$

$$\because t^2/2 = P$$

$$= (1 + P/n)^n$$

applying limit on both sides we have

$$\lim_{n \rightarrow \infty} M_{Z_n}(t) = \lim_{n \rightarrow \infty} (1 + P/n)^n$$

$$= e^P$$

$$P = t^2/2$$

$$= e^{t^2/2} = M_X(t)$$

This indicates that for  $Z_n$  has the mgf as the standard normal distribution

$$M_{Z_n}(t) \rightarrow M_X(t) \text{ when } n \rightarrow \infty$$

Q:- Let  $S_n$  be the number of successes in  $n$  Bernoulli trials with probability  $P = \frac{1}{2}$  for success on each trial. Then show that with the help of Chebyshev inequality for any  $\epsilon > 0$

$$a) P\left[\left|\frac{S_n}{n} - P\right| > \epsilon\right] \leq \frac{P(1-P)}{n\epsilon^2}$$

b) Find max possible value for  $P(1-P)$  if  $P$  lies between 0 and 1 and show that  $P\left[\left|\frac{S_n}{n} - \frac{1}{2}\right| > \epsilon\right] \leq \frac{1}{4n\epsilon^2}$

We know that

$$P \left[ |\bar{x} - E(\bar{x})| > \epsilon \right] \leq \frac{V(\bar{x})}{\epsilon^2}$$

In our case

$$\bar{x} = \frac{S_n}{n}$$

$$P \left[ \left| \frac{S_n}{n} - E\left(\frac{S_n}{n}\right) \right| > \epsilon \right] \leq \frac{V(S_n/n)}{\epsilon^2} \quad \text{--- (1)}$$

$$\begin{aligned} E\left(\frac{S_n}{n}\right) &= \frac{1}{n} E(S_n) = \frac{1}{n} E\{X_1 + X_2 + \dots + X_n\} \\ &= \frac{n\mu}{n} = \mu = P \end{aligned}$$

$$\begin{aligned} V\left(\frac{S_n}{n}\right) &= \frac{1}{n^2} V(X_1 + X_2 + \dots + X_n) \\ &= \frac{1}{n^2} (\sigma^2 + \sigma^2 + \dots + \sigma^2) \\ &= \frac{n\sigma^2}{n^2} = \frac{\sigma^2}{n} = \frac{P(1-P)}{n} \end{aligned}$$

Now (1) becomes As

$$P \left[ \left| \frac{S_n}{n} - P \right| > \epsilon \right] \leq \frac{P(1-P)}{n\epsilon^2} \quad \text{--- (2)}$$

b) Since  $P$  lies between 0 and 1 and we also know that  $P(1-P)$  has its maximum value at  $P = \frac{1}{2}$ . Now using eq (2) we have

$$P \left[ \left| \frac{S_n}{n} - \frac{1}{2} \right| > \epsilon \right] \leq \frac{1}{4n\epsilon^2} \quad \because P(1-P) = \frac{1}{2} \cdot \frac{1}{2} = \frac{1}{4}$$

Q:- Let  $x$  be a random variable with  $E(x)=0$  and  $V(x)=1$ . what integer value of  $k$  will assure us that

$$P[|X| > k] \leq 0.01$$

using Chebyshev inequality

$$P[|x - E(x)| > \epsilon] \leq \frac{V(x)}{\epsilon^2}$$

here  $E(x)=0$  and  $V(x)=1$

$$P[|x - 0| > \epsilon] \leq \frac{1}{\epsilon^2}$$

let for any no  $k$

$$P[|X| > k] \leq \frac{1}{k^2} \text{ --- (1)}$$

it is given that

$$P[|X| > k] \leq 0.01 \text{ --- (2)}$$

Comparing eq (1) and (2)

$$\frac{1}{k^2} = 0.01$$

$$k^2 = 100$$

$$k = 10.$$

Q:- Let  $S_{100}$  be the numbers of heads that turn up in 100 tosses of a fair coin. Use Central limit theorem to estimate probability of

a)  $P(S_{100} < 45) = ?$

b)  $P(45 \leq S_{100} < 55) = ?$

c)  $P(S_{100} > 63) = ?$       d)  $P(S_{100} < 47) = ?$

we have

$$Z_n = \frac{S_n - np}{\sqrt{npq}}$$

Here  $np = 100 \times \frac{1}{2} = 50$

$$\sqrt{npq} = \sqrt{100 \times \frac{1}{2} \times \frac{1}{2}} = 5$$

Now

a)  $P(S_{100} < 45) = P\left(\frac{S_{100} - 50}{5} < \frac{45 - 50}{5}\right)$

$$P(Z_{100} < -1) = P[Z < -1]$$

$$= 0.15866$$

b)  $P(45 \leq S_{100} < 55)$

$$= P\left[\frac{45 - 50}{5} \leq \frac{S_{100} - 50}{5} \leq \frac{55 - 50}{5}\right]$$

$$= P[-1 \leq Z_{100} < 1]$$



$$= P[0 \leq z \leq 1] - P[0 \leq z \leq -1]$$

$$= 0.84134 - 0.15866 = 0.68268$$

using standard normal probabilities

$$c) P(S_{100} > 63) = P\left[\frac{S_{100} - S_0}{5} > \frac{63 - S_0}{5}\right]$$

$$P[Z_{100} > 2.6] = P[Z > 2.6]$$

$$= 1 - P[0 \leq z \leq 2.6]$$

$$= 1 - 0.99534 = 0.00466$$

$$d) P(S_{100} < 47)$$

$$= P\left(\frac{S_{100} - S_0}{5} < \frac{47 - S_0}{5}\right)$$

$$= P[Z_{100} < -0.6]$$

$$= P[Z < -0.6] = 0.27425$$

Q:- There are 100 men on a plane. Let  $X_i$  be the weight in pounds of  $i$ th man on the plane. Suppose that  $X_i$ 's are independent and identically distributed and  $E(X_i) = \mu = 170$  and  $\sigma_{X_i} = \sigma = 30$ . Find the probability that total weight of the men on the plane exceeds 18000 pounds.

Let  $w = X_1 + X_2 + \dots + X_{100}$   
= weight of all persons.

$$E(X_i) = \mu = 170$$

$$\sigma_{X_i} = 30$$

$$P(w > 18000) = ?$$

$$\text{Now } E(w) = n\mu = n \times 170 = 100 \times 170 \\ = 17000$$

$$V(w) = n\sigma^2 = 100 \times (30^2) = 90,000$$

$$\sigma_w = 300$$

$$P(w > 18000) = P\left(\frac{w - 17000}{300} > \frac{18000 - 17000}{300}\right)$$

$$= P(Z > 3.33)$$

$$= 1 - P(0 \leq Z \leq 3.33)$$

$$= 1 - 0.99957 = 0.00043$$

using standard area probabilities

Q:- Let  $X_1, X_2, \dots, X_{25}$  be independent identically distributed with following pmf

$$P(X) = 0.6 \quad \text{if } X = 1$$

$$P(X) = 0.4 \quad \text{if } X = -1$$

and let  $Y = X_1 + X_2 + \dots + X_n$

using Central limit theorem with

Continuity Correction Factor estimate  
probability of  $Y$  lies b/w 4 and 6.

∴ In Probability theory, a Continuity Correction is an adjustment that is made when a discrete distribution is approximated by a Continuous distribution.

جب (Discrete distribution) سے (Continuous distribution) میں جاتے ہیں

تو (D) اور (C) کے فرق یا (Gap) کو ختم کرنے کے  
لئے Continuity correction (use) کرتے ہیں۔

we have to calculate

$$P(4 \leq Y \leq 6)$$

using Continuity Correction.

$$P\{3.5 \leq Y \leq 6.5\} = ?$$

For this we need mean and variance  
of  $Y$ .

$$\begin{aligned} E(X_i) &= E(X_1) + E(X_2) \\ &= X_1 P(X_1) + X_2 P(X_2) \\ &= 0.6 - 0.4 = 0.2 \end{aligned}$$

$$E(x_i^2) = \sum x_i^2 P(x)$$

$$= 0.6 + 0.4 = 1$$

$$V(x) = E(x_i^2) - [E(x_i)]^2$$

$$= 1 - (0.2)^2 = 0.96$$

Now

$$Y = X_1 + X_2 + \dots + X_{25}$$

$$E(Y) = 25\mu = 25(0.2) = 5$$

$$V(Y) = V(X_1 + X_2 + \dots + X_{25})$$

$$= 25\sigma^2 = 25(0.96) = 24$$

$$= P\{3.5 \leq Y \leq 6.5\}$$

$$= P\left\{\frac{3.5-5}{(24)^{1/2}} \leq \frac{Y-5}{\sqrt{24}} \leq \frac{6.5-5}{\sqrt{24}}\right\}$$

$$= P\left\{\frac{3.5-5}{4.89} \leq \frac{Y-5}{4.89} \leq \frac{6.5-5}{4.89}\right\}$$

$$= P\{-0.31 \leq Z \leq 0.31\}$$

$$= 0.62172 - 0.37828$$

$$= 0.24344$$

Q :- You have invited 64 guest to a party. You need to make sandwiches for the guest, you believe that guest might need 0, 1 or 2 sandwiches with

probability  $\frac{1}{4}$ ,  $\frac{1}{2}$  and  $\frac{1}{4}$ . You assume that the number of sandwiches each guest need is independent from other guest. How many sandwiches should you make so that you are 95% sure that there is no shortage.

Let  $X_i$  be the no. of sandwiches that the  $i^{\text{th}}$  person need and let  $Y = X_1 + X_2 + \dots + X_4$ . we need  $Y$  such that

$$P\{Y \leq y\} \geq 0.95$$

$$E(X_i) = \sum X_i P(X_i)$$

$$= X_1 P(X_1) + X_2 P(X_2) + X_3 P(X_3)$$

$$= 0 + \frac{1}{2} + 2 \cdot \frac{1}{4} = 0 + \frac{1}{2} + \frac{1}{2} = 1$$

$$E(X_i^2) = \sum X_i^2 P(X_i)$$

$$= 0 + \frac{1}{2} + 4 \cdot \frac{1}{4} = \frac{3}{2} = 1.5$$

$$V(X_i) = 1.5 - (1)^2 = 0.5$$

$$P\{Y \leq y\} \geq 0.95$$

$$P\left\{\frac{Y - E(Y)}{\sigma_Y} \leq \frac{y - E(Y)}{\sigma_Y}\right\} \geq 0.95 \quad \text{--- (1)}$$

$$E(Y) = E(X_1 + X_2 + \dots + X_{64})$$

$$= 64\mu = 64 \times 1 = 64$$

$$V(Y) = 64\sigma^2 = 64 \times 0.5 = 32.$$

Now (1) becomes As.

$$P\left[Z \leq \frac{y-64}{32}\right] \geq 0.95 \quad \sim$$

$$P\left[Z \leq \frac{y-64}{32}\right] = 1.645$$

$$\frac{y-64}{32} = 1.645$$

$$y = (1.645 \times 32) + 64$$

$$y = 74.$$

Q:- Let  $X_1, X_2, \dots, X_n$  be independent and identically distributed random variable follow exponential distribution with parameter  $\lambda$  and  $\lambda=1$ . Let  $\bar{X} = \frac{X_1 + X_2 + \dots + X_n}{n}$  how large should  $n$  be such that  $P(0.9 < \bar{X} < 1.1) \geq 0.95$

Since  $X \sim \text{exp}(1)$

$$E(X_i) = \frac{1}{\lambda} = 1$$

$$\text{var}(X_i) = \frac{1}{\lambda^2} = 1$$

$$P[0.9 < \bar{x} < 1.1] \geq 0.95$$

$$\text{let } Y = X_1 + X_2 + \dots + X_n$$

$$P\left[0.9 < \frac{Y}{n} < 1.1\right] \geq 0.95 \quad \text{--- (1)}$$

$$\text{Now } E(Y) = E(\sum X_i) = \sum \{E(X_i)\}$$

$$= \sum (1) = n$$

$$\text{var}(Y) = \text{var}(\sum X_i) = \sum \{\text{var}(X_i)\}$$

$$= \sum (1) = n$$

$$\text{Now (1)} \Rightarrow P[0.9n < Y < 1.1n] \geq 0.95$$

$$P\left[\frac{0.9n - n}{\sqrt{n}} < Z_Y < \frac{1.1n - n}{\sqrt{n}}\right] \geq 0.95$$

$$P\left[\frac{n(0.9 - 1)}{\sqrt{n}} < Z_Y < \frac{n(1.1 - 1)}{\sqrt{n}}\right] \geq 0.95$$

$$P[-0.1\sqrt{n} < Z_Y < 0.1\sqrt{n}] \geq 0.95$$

According to probability Rule

$$\Phi(0.1\sqrt{n}) - \Phi(-0.1\sqrt{n}) \geq 0.95$$

$$\because \Phi(z) = 1 - \Phi(-z)$$

$$\Phi(0.1\sqrt{n}) - 1 + \Phi(0.1\sqrt{n}) \geq 0.95$$

$$2\Phi(0.1\sqrt{n}) - 1 \geq 0.95$$

$$2\Phi(0.1\sqrt{n}) \geq 1.95$$

$$\Phi(0.1\sqrt{n}) \geq 1.95/2$$

$$\Phi(0.1\sqrt{n}) \geq 0.975$$

$$0.1\sqrt{n} \geq \Phi^{-1}(0.975)$$

$$0.1\sqrt{n} \geq 1.96$$

$$\sqrt{n} \geq \left( \frac{1.96}{0.1} \right) \Rightarrow n \geq \left( \frac{1.96}{0.1} \right)^2$$

$$n \geq 384.16.$$



Markov chain :- A Markov chain is a mathematical system that experiences transition from one state to another according to some probabilistic rules. The defining characteristic of a Markov chain is that no matter how the process arrived at its present state, the possible future states are fixed. In other words, the probability of transitioning to any particular state is dependent solely on the current state and time elapsed.

Transition Matrices :- A transition matrix  $P_t$  for Markov chain at time  $t$  is a matrix containing information on the probability of transitioning between states. In particular, given an ordering of a matrix's rows and columns by state spaces. This means each row of a matrix is a probability vector, and the sum of its entries is 1.

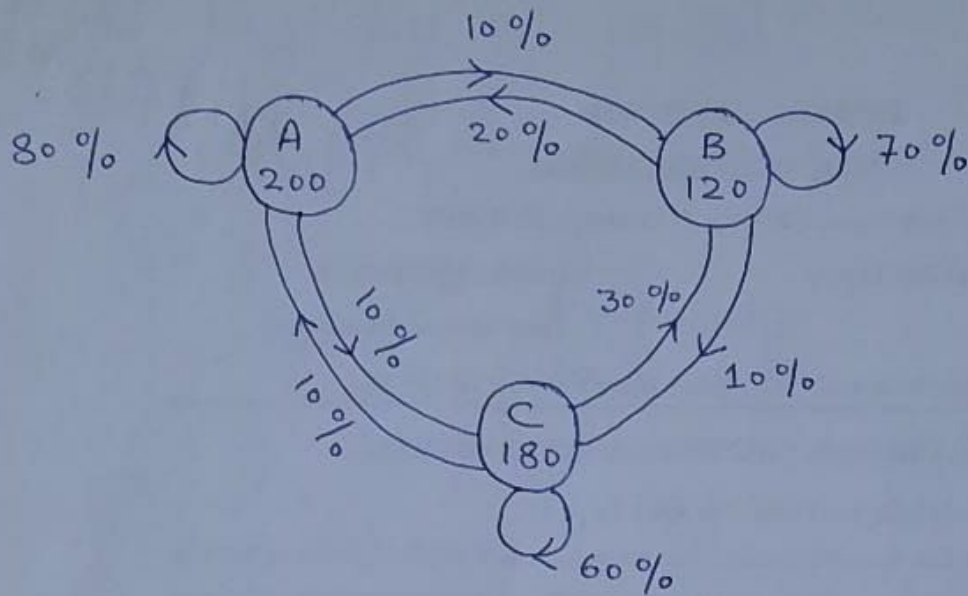
Most of our study of probability has dealt with independent trial process. When we study a sequence of chance experiments from an independent trials process, the possible outcomes of each experiment are the same and occur with the same probability. Further, knowledge of the outcomes of the previous experiments does not influence our predictions for the outcomes of the next experiment.

Modern probability theory studies chance process for which the knowledge of previous outcomes influences predictions for future experiments. In principle, when we observe a sequence of chance experiments, all of the past outcomes could influence our predictions for the next experiments. For example, this should be the case in predicting a student's grades on a sequence of exams in a course. But to allow this much generally would make it very difficult to prove general results.

In 1907, Markov began the study of an important new type of chance process. In this process, the outcome of a given experiment can effect the outcome of the next experiment. This type of process is called a Markov chain.

Question :- These are three shopping malls in a certain colony. A sample of 500 people are collected from these malls on the basis of their choices about these shopping malls. From the data given below in the form of chain <sup>calculate the</sup> Next state to illustrate the condition of these shopping malls.

or transition diagram



$$[\text{NEXT STATE}] = \begin{bmatrix} \text{Matrix of} \\ \text{Transition} \\ \text{Probabilities} \end{bmatrix} [\text{Current state}]$$

$$\begin{bmatrix} A_1 \\ B_1 \\ C_1 \end{bmatrix} = \begin{matrix} A \\ B \\ C \end{matrix} \begin{bmatrix} A & B & C \\ 0.8 & 0.2 & 0.1 \\ 0.1 & 0.7 & 0.3 \\ 0.1 & 0.1 & 0.6 \end{bmatrix} \begin{bmatrix} 200 \\ 120 \\ 180 \end{bmatrix} \begin{bmatrix} 200/500 \\ 120/500 \\ 180/500 \end{bmatrix}$$

OR

$$= \begin{bmatrix} (0.8)(0.4) + (0.2)(0.4) + (0.1)(0.36) \\ (0.1)(0.4) + (0.7)(0.24) + (0.3)(0.36) \\ (0.1)(0.4) + (0.1)(0.24) + (0.6)(0.36) \end{bmatrix}$$

$$\Rightarrow \begin{bmatrix} 0.404 \\ 0.316 \\ 0.280 \end{bmatrix} \rightarrow \text{Sum should be one}$$

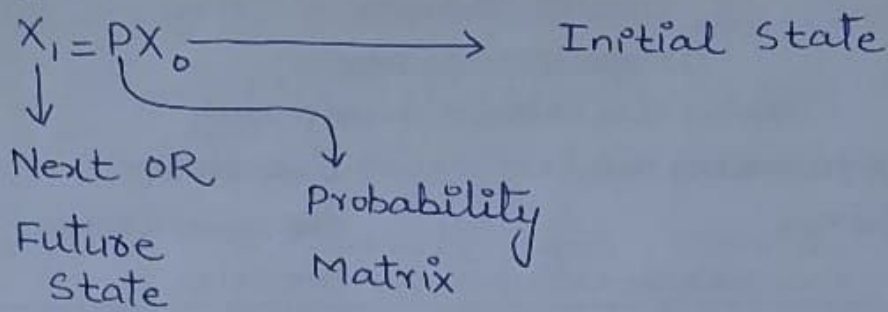
$$0.404 + 0.316 + 0.280 = 1$$

Now the Next state is as follow

$$\begin{bmatrix} 0.404 \times 500 = 202 \\ 0.316 \times 500 = 158 \\ 0.280 \times 500 = 140 \end{bmatrix} \quad \begin{matrix} \text{Total Sampled people} \\ 202 + 158 + 140 = 500 \end{matrix}$$



why they call it chains

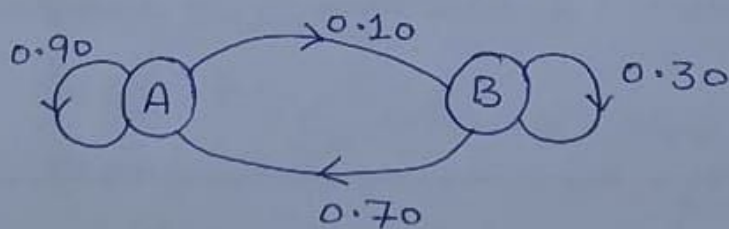


$$X_2 = PX_1 \quad X_3 = PX_2 \quad \text{and so on.}$$

Question :- Suppose that an orange juice company controls 20% of the orange juice market. ~~Suppose~~ They hire a market research company to predict the effects on an aggressive ad campaign. Suppose they conclude

- Someone using Brand A will stay with Brand A with 90% probability, after 1 week.
- Someone Not using Brand A will switch to Brand A with 70% probability, after 1 week.

Transition Diagram



$$\begin{bmatrix} \text{Next} \\ \text{state} \end{bmatrix} = \begin{bmatrix} \text{Transition} \\ \text{Probability} \\ \text{Matrix} \end{bmatrix} \begin{bmatrix} \text{Current state} \end{bmatrix}$$