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 Show that $\hat{\theta}$ is consistent but $\text{bias}(\hat{\theta}) \neq 0$
 i.e. $\text{bias}(\hat{\theta}) \rightarrow 0$ as $n \rightarrow \infty$.

Sol As bias is not zero then MSE can't be zero

$$\text{Bias}(\hat{\theta}) = E(\hat{\theta}) - \theta$$

$$E(\hat{\theta}) = \sum_{\hat{\theta} = \theta, \theta+n} \hat{\theta} P(\hat{\theta})$$

$$= \theta P(\hat{\theta} = \theta) + (\theta+n) P(\hat{\theta} = \theta+n)$$

$$= \theta \left(\frac{n-1}{n}\right) + (\theta+n) \frac{1}{n}$$

$$= \frac{n\theta - \theta}{n} + \frac{\theta}{n} + 1$$

$$E(\hat{\theta}) = \theta + 1$$

So $\text{Bias}(\hat{\theta}) = E(\hat{\theta}) - \theta = \theta + 1 - \theta = 1$
 which is constant number and constant never approaches to zero as $n \rightarrow \infty$, so

$$\text{bias}(\hat{\theta}) \neq 0$$

$$E(\hat{\theta}^2) = \theta^2 \left(\frac{n-1}{n}\right) + (\theta+n)^2 \frac{1}{n}$$

$$= \theta^2 - \frac{\theta^2}{n} + \frac{\theta^2}{n} + n + 2\theta$$

$$= \theta^2 + 2\theta + n$$

$$V(\hat{\theta}) = E(\hat{\theta}^2) - (E(\hat{\theta}))^2$$

$$= \theta^2 + 2\theta + n - (\theta+1)^2$$

$$= \theta^2 + 2\theta + n - \theta^2 - 1 - 2\theta = n - 1$$

$$\begin{aligned}
 E(\hat{\theta}_n^2) &= \int_0^{\theta} y^2 \cdot \frac{ny^{n-1}}{\theta^n} dy \\
 &= \frac{n}{\theta^n} \int_0^{\theta} y^{n+1} dy = \frac{n}{\theta^n} \left(\frac{y^{n+2}}{n+2} \Big|_0^{\theta} \right) \\
 &= \frac{n}{\theta^n} \frac{\theta^{n+2}}{n+2} = \frac{n}{n+2} \theta^{n+2-n}
 \end{aligned}$$

$$E(\hat{\theta}_n) = \frac{n}{n+2} \theta^2$$

Now from (2)

$$\begin{aligned}
 \text{Var}(\hat{\theta}_n) &= E(\hat{\theta}_n^2) - (E(\hat{\theta}_n))^2 \\
 &= \frac{n}{n+2} \theta^2 - \left(\frac{n\theta}{n+1} \right)^2 \\
 &= \frac{n\theta^2}{n+2} - \frac{n^2\theta^2}{(n+1)^2}
 \end{aligned}$$

$$\text{Var}(\hat{\theta}_n) = \frac{n\theta^2(n+1)^2 - n^2\theta^2(n+2)}{(n+2)(n+1)^2}$$

$$= \frac{n\theta^2(n^2+1+2n) - n^3\theta^2 - 2n^2\theta^2}{(n+2)(n+1)^2}$$

$$= \frac{\cancel{n^3}\theta^2 + n\theta^2 + 2n^2\theta^2 - \cancel{n^3}\theta^2 - 2n^2\theta^2}{(n+2)(n+1)^2}$$

$$\text{Var}(\hat{\theta}_n) = \frac{n\theta^2}{(n+2)(n+1)^2}$$

Lecture 6: 22-6-2020

Problem:- Let X_1, \dots, X_n be a r.v.s from a $\text{Uni}(0, \theta)$ where θ is unknown. Define the estimator.

$$\hat{\theta}_n = \max\{X_1, X_2, \dots, X_n\}$$

(a) Find bias of $\hat{\theta}_n$ $B(\hat{\theta}_n)$

(b) Find the MSE of $\hat{\theta}_n$ $MSE(\hat{\theta}_n)$

(c) Is $\hat{\theta}_n$ a consistent estimator of θ ?

Sol If $X \sim \text{Uni}(0, \theta)$, then the PDF and CDF of X are given by

$$f_X(x) = \begin{cases} \frac{1}{\theta} & 0 \leq x \leq \theta \\ 0 & \text{o.w} \end{cases}$$

and

$$F_X(x) = \begin{cases} 0 & x < 0 \\ \frac{x}{\theta} & 0 \leq x \leq \theta \end{cases}$$

the pdf of $\hat{\theta}_n$ is given by

$$\begin{aligned} f_{\hat{\theta}_n}(y) &= n f_X(y) [F_X(y)]^{n-1} \\ &= n \frac{1}{\theta} \left(\frac{y}{\theta}\right)^{n-1} = n \left(\frac{1}{\theta}\right) \frac{y^{n-1}}{\theta^{n-1}} \\ &= \begin{cases} \frac{n y^{n-1}}{\theta^n} & 0 \leq y \leq \theta \end{cases} \end{aligned}$$

by direct method

→ By definition

$$= P\{|\hat{\theta} - \theta| > \epsilon\} = P\left(\left|\frac{\bar{X}}{4} - \theta\right| > \epsilon\right)$$

$$= P\left\{\left|\frac{\bar{X}/4 - \theta}{\sqrt{\bar{X}/4}}\right| > \frac{\epsilon}{\sqrt{\bar{X}/4}}\right\}$$

$$= P\left\{\frac{|\bar{X}/4 - \theta|}{\sigma^2/4n} > \frac{\epsilon}{\sigma^2/4n}\right\}$$

$$= P\left\{|Z| > \frac{4n\epsilon}{\sigma^2}\right\}$$

$$= 1 - P\left\{|Z| \leq \frac{4n\epsilon}{\sigma^2}\right\}$$

$$\lim_{n \rightarrow \infty} P\{|\hat{\theta} - \theta| > \epsilon\} = 1 - P(|Z| \leq \infty)$$

$$= 1 - P(-\infty < Z < \infty)$$

$$= 1 - \int_{-\infty}^{\infty} f(z) dz = 1 - 1 = 0$$

hence proved $P\{|\hat{\theta} - \theta| > \epsilon\} = 0$ consistent.

Example: Suppose $\hat{\theta}$ is an estimator of θ with prob function

$$P(\hat{\theta} = \theta) = \frac{n-1}{n}$$

$$P(\hat{\theta} = \theta + n) = 1/n$$

and no other value of θ are possible.

A consistent estimate has insignificant errors
as sample size increases. More specifically
the more data to collect, a consistent
estimator will be close to the real

popⁿ parameter you are trying to measure.

The sample mean and sample variance
are two well known consistent estimators.

Let $\hat{\theta}$ be an estimator of θ based on a
sample of size n . Then $\hat{\theta}$ is a consistent
estimator of θ , if $\text{var}(\hat{\theta}) \rightarrow 0$ as $n \rightarrow \infty$

To illustrate, let us consider \bar{X} , the sample mean
based on random of size n . We know that

$$E(\bar{X}) = \mu, \quad \text{var}(\bar{X}) = \frac{\sigma^2}{n}$$

Now as $n \rightarrow \infty$ $\frac{\sigma^2}{n} \rightarrow 0$

$$\lim_{n \rightarrow \infty} \text{var}(\bar{X}) = \lim_{n \rightarrow \infty} \frac{\sigma^2}{n} = 0$$

Hence \bar{X} is an U/B estimator of μ .

Now from (1),

$$\begin{aligned}
 \text{MSE}(\hat{\theta}_n) &= \text{var}(\hat{\theta}_n) + (B(\hat{\theta}_n))^2 \\
 &= \frac{n\theta^2}{(n+2)(n+1)^2} + \left(\frac{-\theta}{n+1}\right)^2 \\
 &= \frac{n\theta^2}{(n+2)(n+1)^2} + \frac{\theta^2}{(n+1)^2} \\
 &= \frac{n\theta^2 + \theta^2(n+2)}{(n+2)(n+1)^2} \\
 &= \frac{n\theta^2 + \theta^2 n + 2\theta^2}{(n+2)(n+1)^2} \\
 &= \frac{2n\theta^2 + 2\theta^2}{(n+2)(n+1)^2}
 \end{aligned}$$

$$\text{MSE}(\hat{\theta}_n) = \frac{2\theta^2(n+1)}{(n+2)(n+1)^2} = \frac{2\theta^2}{(n+2)(n+1)}$$

$$\text{MSE}(\hat{\theta}_n) = \frac{2\theta^2}{(n+2)(n+1)}$$

(c) Is $\hat{\theta}_n$ is consistent estimator of θ .

$$\lim_{n \rightarrow \infty} \text{MSE}(\hat{\theta}_n) = \lim_{n \rightarrow \infty} \frac{2\theta^2}{(n+2)(n+1)} = 0$$

$\hat{\theta}_n$ is a consistent estimator of θ .

Also for proposition, $\hat{p} = \frac{X}{n}$

$$E(\hat{p}) = E\left(\frac{X}{n}\right) = \frac{E(X)}{n} = \frac{np}{n} = p$$

So \hat{p} is unbiased estimator of p

$$\text{Now } \text{var}(\hat{p}) = \text{var}\left(\frac{X}{n}\right) = \frac{1}{n^2} \text{var}(X)$$

$$\text{var}(\hat{p}) = \frac{1}{n^2} (npq) = \frac{pq}{n}$$

$$\lim_{n \rightarrow \infty} \text{var}(\hat{p}) = \lim_{n \rightarrow \infty} \frac{pq}{n} = 0$$

hence \hat{p} is a consistent estimator

of p .

Example (Lecture 5, page 3)

Investigate $\hat{\theta}$ is equal to $\bar{X}/4$ i.e. $\hat{\theta} = \bar{X}/4$
is an unbiased estimator and consistent.

$$E(\hat{\theta}) = \theta \quad \text{where } \hat{\theta} = \bar{X}/4$$

Sol

$$\text{var}(\hat{\theta}) = \frac{1}{16} \text{var}(\bar{X})$$

$$V(\bar{X}) = \frac{\sigma^2}{n}$$

$$= \frac{1}{16} \frac{\sigma^2}{n} = \frac{1}{16} \frac{40^2}{n}$$

$$\sigma^2 = 40^2$$

$$\text{var}(\hat{\theta}) = \frac{40^2}{4n}$$

$$\rightarrow \text{now } \lim_{n \rightarrow \infty} \text{var}(\hat{\theta}) = \lim_{n \rightarrow \infty} \frac{40^2}{4n} = 0$$

hence proved

To find the bias we have

$$B(\hat{\theta}_n) = E[\hat{\theta}_n] - \theta$$

Now find

$$E[\hat{\theta}_n]$$

$$E(\hat{\theta}_n) = \int_0^{\theta} y \cdot \frac{ny^{n-1}}{\theta^n} dy$$

$$= \frac{n}{\theta^n} \int_0^{\theta} y^n dy = \frac{n}{\theta^n} \left(\frac{y^{n+1}}{n+1} \Big|_0^{\theta} \right)$$

$$E(\hat{\theta}_n) = \frac{n}{\theta^n} \left(\frac{\theta^{n+1}}{n+1} \right) = \frac{n}{n+1} \theta^{n+1-n} = \frac{n}{n+1} \theta$$

Now find bias

$$B(\hat{\theta}_n) = \frac{n}{n+1} \theta - \theta = \frac{n\theta - n\theta - \theta}{n+1} = \frac{-\theta}{n+1}$$

(b) To find $MSE(\hat{\theta}_n)$, we can write

$$MSE(\hat{\theta}_n) = \text{Var}(\hat{\theta}_n) + (B(\hat{\theta}_n))^2$$

$$MSE(\hat{\theta}_n) = \text{Var}(\hat{\theta}_n) + \frac{\theta^2}{(n+1)^2} \quad \text{--- (1)}$$

Thus

$$\text{Var}(\hat{\theta}_n) = E[\hat{\theta}_n^2] - (E(\hat{\theta}_n))^2 \quad \text{--- (2)}$$

$E(\hat{\theta}_n^2)$ will be solve in similar way as $E(\hat{\theta}_n)$