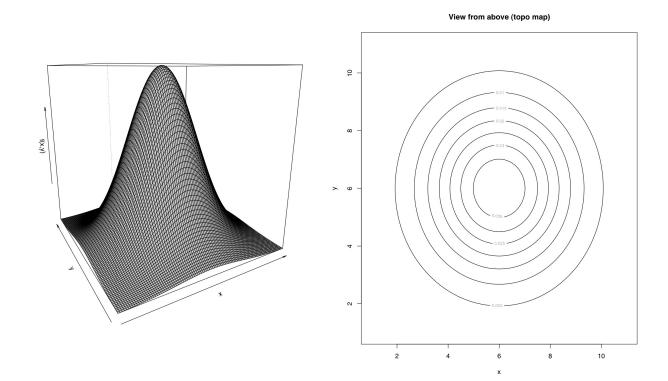
Chapter 5: JOINT PROBABILITY <u>DISTRIBUTIONS</u>

Part 3: The Bivariate Normal

Section 5.5.2

Linear Functions of Random Variables

Section 5.6



The bivariate normal is kind of nifty because...

- The marginal distributions of X and Y are both univariate normal distributions.
- The conditional distribution of Y given X is a normal distribution.
- The conditional distribution of X given Y is a normal distribution.
- Linear combinations of X and Y (such as Z = 2X + 4Y) follow a normal distribution.
- It's normal almost any way you slice it.

• Bivariate Normal Probability Density Function

The parameters: $\mu_X, \mu_Y, \sigma_X, \sigma_Y, \rho$ $f_{XY}(x,y) = \frac{1}{2\pi\sigma_X\sigma_Y\sqrt{(1-\rho^2)}} \times \exp\left\{\frac{-1}{2(1-\rho^2)}\left[\frac{(x-\mu_X)^2}{\sigma_X^2} - \frac{2\rho(x-\mu_X)(y-\mu_Y)}{\sigma_X\sigma_Y} + \frac{(y-\mu_Y)^2}{\sigma_Y^2}\right]\right\}$

for $-\infty < x < \infty$ and $-\infty < y < \infty$, with

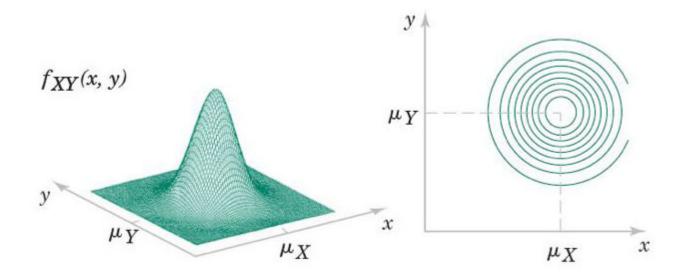
parameters
$$\sigma_X > 0$$
, $\sigma_Y > 0$,
 $-\infty < \mu_X < \infty$, $-\infty < \mu_Y < \infty$,
and $-1 < \rho < 1$.

Where ρ is the correlation between X and Y.

The other parameters are the needed parameters for the marginal distributions of X and Y.

• Bivariate Normal

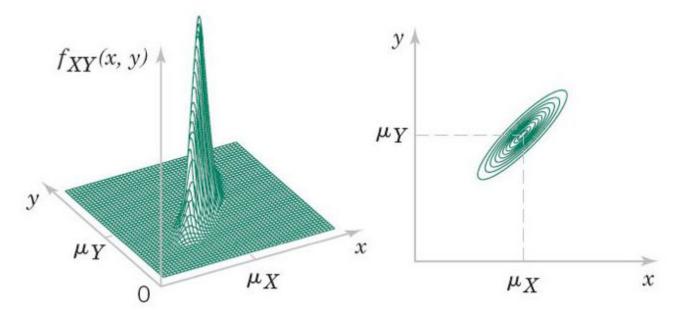
When X and Y are **independent**, the contour plot of the joint distribution looks like concentric circles (or ellipses, if they have different variances) with major/minor axes that are parallel/perpendicular to the x-axis:



The center of each circle or ellipse is at (μ_X, μ_Y) .

• Bivariate Normal

When X and Y are **dependent**, the contour plot of the joint distribution looks like concentric diagonal ellipses, or concentric ellipses with major/minor axes that are NOT parallel/perpendicular to the x-axis:



The center of each ellipse is at (μ_X, μ_Y) .

• Marginal distributions of X and Y in the Bivariate Normal

Marginal distributions of X and Y are normal:

$$X \sim N(\mu_X, \sigma_X^2)$$
 and $Y \sim N(\mu_Y, \sigma_Y^2)$

Know how to take the parameters from the bivariate normal and calculate probabilities in a univariate X or Y problem.

• Conditional distribution of Y|x in the Bivariate Normal

The conditional distribution of Y|x is also normal:

$$Y|x \sim N(\mu_{Y|x}, \sigma_{Y|x}^2)$$

$$Y|x \sim N(\mu_{Y|x}, \sigma_{Y|x}^2)$$

where the "mean of Y|x" or $\mu_{Y|x}$ depends on the given *x*-value as

$$\mu_{Y|x} = \mu_Y + \rho \frac{\sigma_Y}{\sigma_X} (x - \mu_X)$$

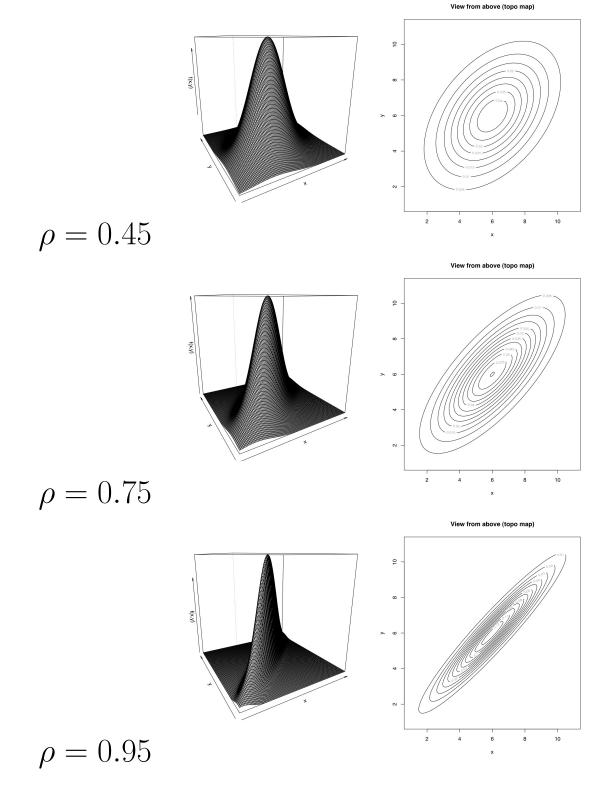
and "variance of Y|x " or $\sigma_{Y|x}^2$ depends on the correlation as

$$\sigma_{Y|x}^2 = \sigma_Y^2 (1 - \rho^2).$$

Know how to take the parameters from the bivariate normal and get a conditional distribution for a given x-value, and then calculate probabilities for the conditional distribution of Y|x (which is a univariate distribution).

Remember that probabilities in the normal case will be found using the z-table.

Notice what happens to the joint distribution (and conditional) as ρ gets closer to +1:



As a last note on the bivariate normal...

Though $\rho = 0$ does not mean X and Y are independent in <u>all cases</u>, for the bivariate normal, this does hold.

For the Bivariate Normal, Zero Correlation Implies Independence

If X and Y have a bivariate normal distribution (so, we know the shape of the joint distribution), then with $\rho = 0$, we have X and Y as independent. • **Example**: From book problem 5-54.

Assume X and Y have a bivariate normal distribution with.. $\mu_X = 120, \sigma_X = 5$ $\mu_Y = 100, \sigma_Y = 2$ $\rho = 0.6$

Determine: (i) Marginal probability distribution of X.

(*ii*) Conditional probability distribution of Y given that X = 125.

Linear Functions of Random Variables Section 5.6

• Linear Combination

Given random variables X_1, X_2, \ldots, X_p and constants c_1, c_2, \ldots, c_p ,

$$Y = c_1 X_1 + c_2 X_2 + \dots + c_p X_p$$

is a linear combination of X_1, X_2, \ldots, X_p .

• Mean of a Linear Function

If
$$Y = c_1 X_1 + c_2 X_2 + \dots + c_p X_p$$
,

 $E(Y) = c_1 E(X_1) + c_2 E(X_2) + \dots + c_p E(X_p)$

• Variance of a Linear Function

If
$$X_1, X_2, \dots, X_p$$
 are random variables, and
 $Y = c_1 X_1 + c_2 X_2 + \dots + c_p X_p$, then
in general
 $V(Y) = c_1^2 V(X_1) + c_2^2 V(X_2) + \dots + c_p^2 V(X_p)$
 $+ 2 \sum_{i < j} \sum_{i < j} c_i c_j cov(X_i, X_j)$

In this class, all our linear combinations of random variables will be done with independent random variables.

If
$$X_1, X_2, \dots, X_p$$
 are **independent**, then
 $V(Y) = c_1^2 V(X_1) + c_2^2 V(X_2) + \dots + c_p^2 V(X_p)$

The most common mistake for finding the variance of a linear combination is to forget to square the coefficients.

• **Example**: Semiconductor product (example 5.26)

A semiconductor product consists of three layers. The variance of the thickness of the first, second, third layers are 25, 40, and 30 nanometers².

What is the variance of the thickness of the final product?

ANS:

• Mean and Variance of an Average

Suppose we randomly generate p observations from the a distribution with mean μ .

Thus, $E(X_i) = \mu$ for i = 1, 2, ..., p. Let $\bar{X} = \frac{(X_1 + X_2 + \dots + X_p)}{p}$ $= \frac{1}{p}X_1 + \frac{1}{p}X_2 + \dots + \frac{1}{p}X_p$

 \overline{X} is as mean and it is a <u>linear combination</u> of the *p* random variable we observed. Because $E(X_i) = \mu$ for $i = 1, 2, \ldots, p$ we have

$$E(\bar{X}) = \mu.$$

 \Rightarrow The expected value of the average of p random variables, all with the same mean μ , is just μ again.

If X_1, X_2, \ldots, X_p are also independent and all with the same variance or $V(X_i) = \sigma^2$ then

$$V(\bar{X}) = \frac{\sigma^2}{p}$$

 \Rightarrow The variance of the average of p identical random variables (i.e. $\frac{\sigma^2}{p}$) is smaller than the variance of a single random variable (i.e. σ^2).

• Reproductive Property of the Normal Distribution

If X_1, X_2, \ldots, X_p are independent, normal <u>random variables</u> with $\overline{E(X_i)} = \mu_i$ and $V(X_i) = \sigma_i^2$ for $i = 1, 2, \ldots, p$,

$$Y = c_1 X_1 + c_2 X_2 + \dots + c_p X_p$$

is a normal random variable with

$$\mu_Y = E(Y) = c_1 \mu_1 + c_2 \mu_2 + \dots + c_p \mu_p$$

and

$$\sigma_Y^2 = V(Y) = c_1^2 \sigma_1^2 + c_2^2 \sigma_2^2 + \dots + c_p^2 \sigma_p^2$$

i.e. $Y \sim N(\mu_Y, \sigma_Y^2)$ as described above.

A linear combination of normal r.v.'s is also normal.

• **Example**: Manufactured part (see example 5.27)

Let the random variables X_1 and X_2 represent lengths of manufactured parts. Assume that X_1 is normal with $E(X_1) = 2$ cm and standard deviation 0.1 cm and that is X_2 is normal with $E(X_2) = 5$ cm and standard deviation 0.2 cm. We will assume X_1 and X_2 are independent.

Find the probability that $2X_1 + 2X_2 < 14.3$

ANS: (next page)

Let Y represent the random variable (a linear combination).

$$Y = 2X_1 + 2X_2$$

$$\mu_Y = E(Y) = 2(2) + 2(5) = 14$$

$$\sigma_Y^2 = V(Y) = 2^2 V(X_1) + 2^2 V(X_2)$$

$$= 4(0.1^2) + 4(0.2^2)$$

$$= 0.04 + 0.16$$

$$= 0.20$$

Because both X_1 and X_2 were normal r.v.'s, the reproductive property of normal r.v.'s gives us...

$$Y \sim N(\mu_Y, \sigma_Y^2)$$
 or $Y \sim N(14, 0.20)$

$$P(Y > 14.5) = P\left(\frac{Y - \mu_Y}{\sigma_Y} > \frac{14.5 - 14}{\sqrt{0.2}}\right)$$

= $P(Z > 1.12)$
= 0.13

• **Example**: Weights of people

Assume that the weights of individuals are independent and normally distributed with a mean of 160 pounds and a standard deviation of 30 pounds. Suppose that 25 people squeeze into an elevator that is designed to hold 4300 pounds.

a) What is the probability that the load exceeds the design limit?

• SPECIAL CASE:

Reproductive Property of the Normal Distribution for a Random Sample

If the X_1, X_2, \ldots, X_p are each **drawn independently from the same normal** distribution, or by notation $X_i \sim N(\mu, \sigma^2)$ for all i, then

$$\bar{X} \sim N(\mu, \frac{\sigma^2}{p})$$
 for any sample size p .

This results because \overline{X} is a linear combination of normals in this situation.